



Department of Pesticide Regulation



Mary-Ann Warmerdam
Director

MEMORANDUM

Arnold Schwarzenegger
Governor

TO: Chuck Salocks, Senior Toxicologist
Office of Environmental Health Hazard Assessment
Pesticide and Environmental Toxicology Branch

FROM: Shifang Fan, Ph.D.
Environmental Scientist
Environmental Monitoring Branch

Original signed by

Bruce Johnson, Ph.D.
Research Scientist III
Environmental Monitoring Branch

Original signed by

Randy Segawa
Environmental Program Manager
Environmental Monitoring Branch
916-324-4137

Original signed by

DATE: February 23, 2010

SUBJECT: METHODOLOGY TO DETERMINE METHYL BROMIDE USE CAP

Introduction

This memorandum provides a detailed explanation of the method used to determine the monthly use limit (township cap) to control seasonal exposure to methyl bromide. As discussed in the interagency workgroup and the memorandum on methyl bromide regulations development (January 29, 2010 memorandum to Andrews and Verder-Carlos), this method has been revised from the one used to support the earlier regulations.

In 2004, the Department of Pesticide Regulation (DPR) established an ambient air concentration limit of 9 ppb (as a monthly average) to control subchronic exposure for health concerns in areas of heavy agricultural use (Title 3, California Code of Regulations section 6447[h]). DPR and county agricultural commissioners limit the use of methyl bromide to no more than 270,000 pounds in any township (6 x 6 mile area) in any month as the primary means to ensure that the 9 ppb limit is not exceeded. DPR determined this "township cap" of 270,000 pounds using linear regression to correlate monitored methyl bromide ambient air concentrations and reported use intensity (Li, et al., 2005) and calculating a tolerance interval (Johnson and Li, 2003). The tolerance interval was calculated using data from both Air Resource Board (ARB) monitoring in 2000 and 2001 (ARB, 2001a; ARB, 2001b; ARB, 2002a; ARB, 2002b) and Alliance of Methyl Bromide Industry (AMBI) monitoring in 2002 (AMBI, 2003) studies.

Subsequent to the 2004 regulations, the regression and tolerance interval approach (Johnson and Li 2003) was reviewed by Dr. Robert Spear (Spear, 2004), Dr. Shu Geng (Geng, 2004) and



Dr. Neil Willits (Willits, 2004), University of California peer reviewers. Willits identified several problems: (1) residuals not normally distributed, (2) heterodasticity, and (3) spatial correlation between nearby monitoring sites. For issues (1) and (2), since the time of the original effort (Johnson and Li 2003) to relate subchronic exposures to use information, DPR requested ARB to conduct ambient air monitoring of methyl bromide in Ventura County to evaluate the effectiveness of the township cap. In response ARB conducted monitoring in 2005 and 2006 at sites in areas of high methyl bromide use. These methyl bromide monitoring data (ARB, 2006; ARB, 2007) were added to increase sample size. To address issue (3), nearby monitoring sites were aggregated in this data analysis. In addition, the aggregation of neighboring sites may help to smooth the data in order to satisfy (1) and (2). A drawback to aggregating monitoring sites is the potential loss of degrees of freedom in the statistical analysis and consequent failure to extract trends and information from the data.

Statistical Consultation

In 2009, DPR further consulted Willits on this data analysis for three questions:

1. Were there better statistical methods for our data analysis? Instead of grouping monitoring concentration and methyl bromide use data sets, Willits suggested and performed a mixed regression model ANOVA with covariance structure using 34 individual date sets (Appendix 1). This independent data analysis showed a significant effect of methyl bromide use on ambient air concentration ($P=0.0192$).
2. In 2001 Kern County monitoring study, four out of six sites had no reported methyl bromide use in a township area and one has irregular township size. This resulted in only one data set at site CRS being used in this data analysis (Table 1). The ratio of air concentration to methyl bromide use at this site was much higher than the rest of data. Was it legitimately correct to treat this data set as an outlier? Willits looked into three residuals (observed, studentized, and predicted) and performed with t-test, sign test, and signed rank test for each of them (Appendix 1). All results showed that residual distribution was significantly abnormal (all $P<0.04$) and the site CRS in Kern County in 2001 was not an extreme value (all $P >0.4$). Therefore, the CRS data was included in this data analysis.
3. How could we compare the relationships between methyl bromide use and ambient air concentration before and after implementing the monthly use township cap in 2004 with limited data sets? Willits used another mixed model ANCOVA for shifts in slope analysis to identify any significant changes by introducing a factor of period (before or after 2004). The results showed a significant covariate interaction ($P=0.0336$) of methyl bromide use and period (Appendix 1). However, the data after 2004 were only from Ventura County. Considering the geographical and atmospheric complications in addition to limited sample size, it is not correct to simply conclude that the implementation of monthly township use cap

since 2004 has reduced ambient air concentration. Also, the data from Ventura in 2002 was within the same range as in 2005 and 2006. Therefore, this data analysis attempted to combine both data sets before and after 2004.

Methods and Materials

Grouping monitoring sites

The raw data used in this regression analysis includes ARB and AMBI monitoring studies in 2000, 2001, 2002, 2005, and 2006 (ARB, 2001a; ARB, 2001b; ARB, 2002a; ARB, 2002b; AMBI, 2003; ARB, 2006; and ARB, 2007). A total of 30 monitoring sites are shown on the monitoring site maps (Figures 1, 2, and 3). Twenty three of the 30 sites were used for regression analysis. Six sites, LS and TO in Ventura County (Figure 1), ARB, MVS, VSD, and ARV in Kern County (Figure 3b) were excluded because there was no methyl bromide use in a township area during the monitoring period (LS, TO, and ARB were the selected background sites). The other site, MET (Figure 3b), was also excluded because the sections around MET site were irregularly shaped (Li, et al., 2005).

Figure 1. Monitoring sites in 2005-2006 ARB studies

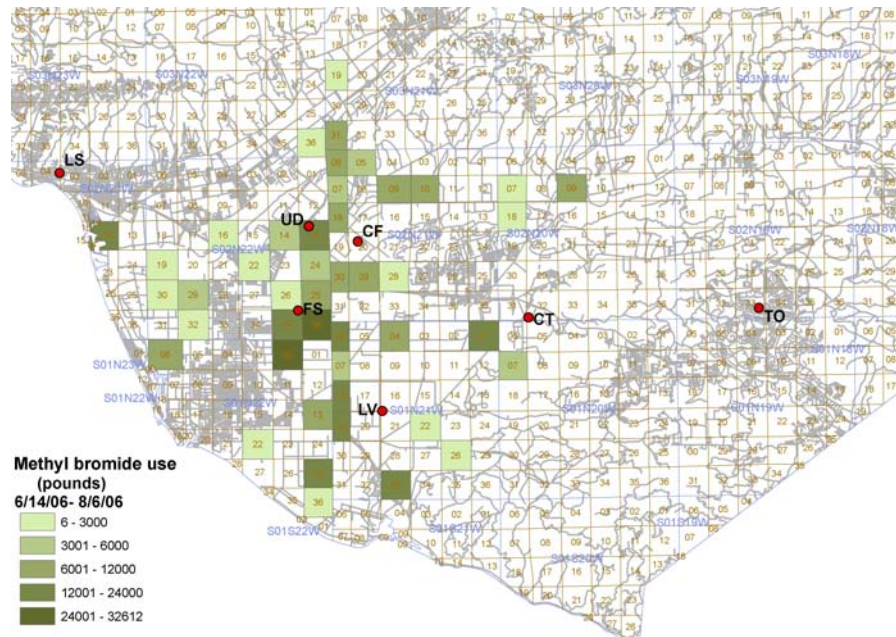


Figure 2. Monitoring sites in 2000-2001 ARB studies

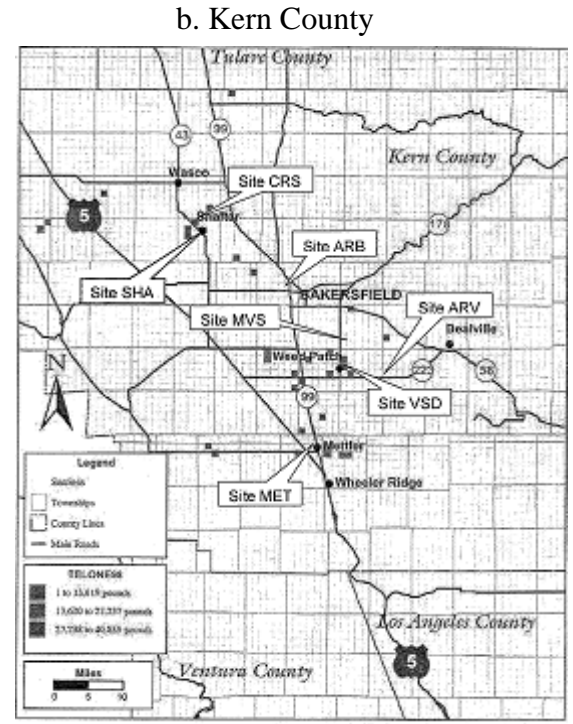
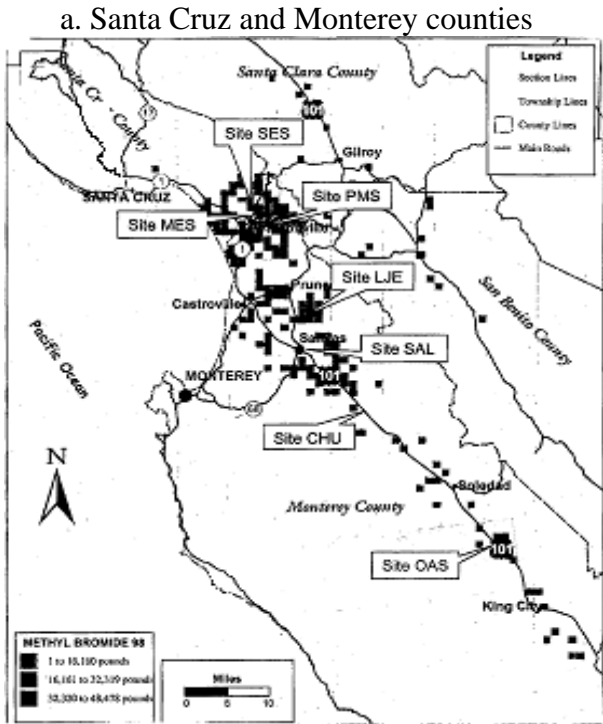


Figure 3. Monitoring sites in 2002 ARBI study

a. Ventura County



From AMBI, 2003

b. Santa Cruz County



Modified from AMBI, 2003

c. Monterey County

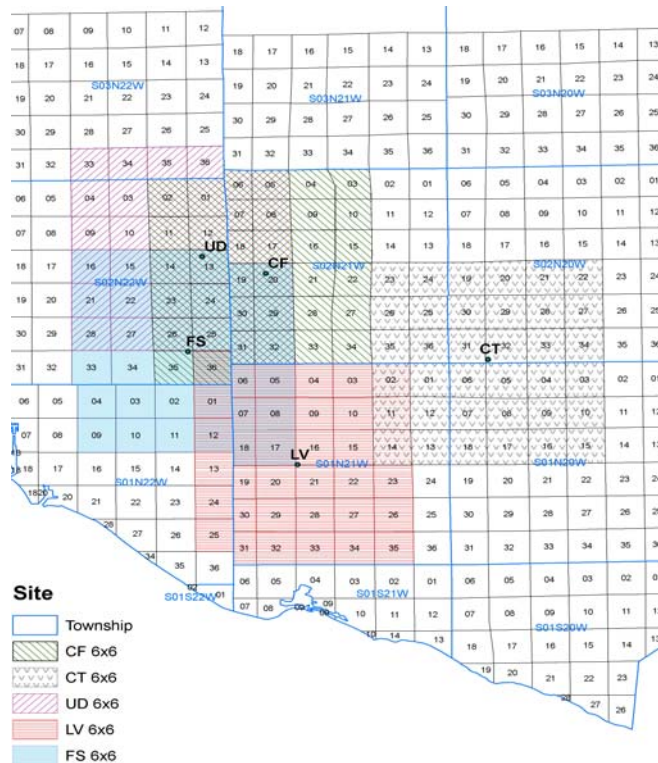


Modified from AMBI, 2003

Data from different years was assumed to be independent. Eleven of the 23 sites were monitored for multiple years. The total sites x years of monitoring resulted in 34 individually measured ambient air concentrations.

Data from different sites were aggregated when a six mile by six mile township-sized area centered on the monitoring site overlapped with an adjacent six mile by six mile township-sized area for a nearby monitoring site (Figure 4). Li et al. (2005) found that the highest correlations between use and ambient air concentrations occurred with domains approximately the size of townships (6x6 mile square areas). The use of the word township in this memorandum refers to surrounding 6x6 square mile area around each monitoring location. The word township can also refer to formal administrative land survey units within California which are used to identify parcels of land. The use of the word township in this memorandum is clear from the context.

Figure 4. Area of 6x6 sections around each monitoring site



The 23 different site locations were aggregated into 9 locations (see number code in Group column in Table 1) based on overlap of the township-sized areas. This resulted in 9 distinct site locations, and 14 ambient air concentrations for the aggregated sites x years data (Table 2).

Ambient air concentration

The methyl bromide air concentration for each group (Table 2) is the average of the average air concentrations at all sites in that group (Table 1). The air concentration at each site is the average of either weekly average air concentrations (Johnson and Li, 2003) or daily monitoring air concentrations over the monitoring period (Fan, 2008). Daily concentrations were 24-hour ambient air samples collected at each monitoring site on four consecutive days a week for six to eight weeks during methyl bromide peak use seasons (ARB, 2001a; ARB, 2001b; ARB 2002a; ARB, 2002b; AMBI, 2003; ARB, 2006; ARB, 2007).

Agricultural use

Methyl bromide use for each group (Table 2) is the average of the normalized monthly use data at all sites in that group (Table 1). The individual site use data were queried from DPR's Pesticide Use Report database (DPR, 2003; DPR, 2008) with the exclusion of outliers (Wilhoit, 1998). The use data in the database is reported by section that contains the field which received the application. A section is usually a square mile. DPR regulates methyl bromide through the use of limits on application within an administrative township based on the California land survey system.

For 2000-2002 data, weekly methyl bromide use was obtained for 5x5 and 7x7 square mile areas centered on the monitoring sites for the monitoring time periods. The use in 6x6 sections around each monitoring site was estimated by the following formula (Johnson and Li, 2003):

$$Use_{(6x6)} = \frac{(36 - 25)}{(49 - 25)} * [Use_{(7x7)} - Use_{(5x5)}] + Use_{(5x5)}$$

where $Use_{(ixi)}$ refers to reported methyl bromide use in areas of (ixi) sections.

The monthly use was normalized by:

$$Use_{monthly} = \frac{30}{7} * Use_{weekly}$$

where Use_{weekly} was the average of the weekly uses during a monitoring period in a township area.

For 2005-2006 data, methyl bromide use was queried in 6x6 mile areas centered on the monitoring site for each entire monitoring period. The monthly use was normalized by:

$$Use_{monthly} = \frac{30}{days(entire_monitoring_period)} * Use_{entire_monitoring_period}$$

Table 1. Monitoring air concentration and monthly use in a township-sized area for each site x year combination, showing the Group codes for grouping as explained in the text

County	Year	Site	Group	Concentration (ppb)	Monthly use (pounds)
Monterey	2000	SAL	2	1.29	39268
Monterey	2000	OAS	3	0.39	5903
Monterey	2000	CHU	4	0.67	13319
Monterey	2000	LJE	2	3.88	66358
Santa cruz	2000	PMS	1	7.73	162024
Santa cruz	2000	SES	1	2.61	80511
Kern	2000	SHA	5	0.79	1876
Kern	2000	CRS	5	2.16	40490
Monterey	2001	SAL	2	1.41	39231
Santa cruz	2001	MES	1	6.14	94270
Monterey	2001	CHU	4	0.58	9883
Monterey	2001	LJE	2	2.86	44597
Santa cruz	2001	PMS	1	3.31	115565
Santa cruz	2001	SES	1	1.14	66660
Kern	2001	CRS	5	2.76	17396
Monterey	2002	MAQ	6	1.12	46163
Monterey	2002	BBC	6	2.08	91768
Santa cruz	2002	WAT	7	3.78	171670
Santa cruz	2002	FRM	7	2.62	150016
Santa cruz	2002	CPW	7	2.06	85698
Ventura	2002	ABD	8	0.76	182374
Ventura	2002	SHA2	8	0.59	61328
Ventura	2002	PVW	8	1.62	155828
Ventura	2002	UWC	8	2.22	195241
Ventura	2005	UD	9	0.29	132800
Ventura	2005	CF	9	0.39	122102
Ventura	2005	FS	9	0.17	175564
Ventura	2005	LV	9	0.22	57560
Ventura	2005	CT	9	0.14	54878
Ventura	2006	UD	9	0.57	51149
Ventura	2006	CF	9	0.52	85545
Ventura	2006	FS	9	0.80	97693
Ventura	2006	LV	9	0.88	73867
Ventura	2006	CT	9	0.45	13716

Table 2. Average air concentration and monthly use after grouping overlapped sites

Site group	County	Year	Concentration (ppb)	Monthly use (pounds)
1	Santa cruz	2000	5.17	121267
1	Santa cruz	2001	3.53	92165
2	Monterey	2000	2.58	52813
2	Monterey	2001	2.14	41914
3	Monterey	2000	0.39	5903
4	Monterey	2000	0.67	13319
4	Monterey	2001	0.58	9883
5	Kern	2000	1.48	21183
5	Kern	2001	2.76	17396
6	Monterey	2002	1.60	68965
7	Santa cruz	2002	2.82	135795
8	Ventura	2002	1.29	148693
9	Ventura	2005	0.24	108581
9	Ventura	2006	0.64	64394

Results and Discussion

Regression Analysis

Using the grouped data sets, a parametric linear regression of the ambient air concentration to the methyl bromide use was attempted:

$$Y = a + bX$$

where *Y* is the methyl bromide air concentration (ppb); *X* is the monthly methyl bromide use (pounds/month) during the monitoring period; *a* and *b* are regression coefficients, intercept and slope, respectively.

Three regression analyses were conducted using Minitab 14 software (Minitab, 2003): A. using all 14 group data sets, B. 3 data sets from Ventura County only, and C. 11 nonVentura County data. The analysis using data sets C was the only one that resulted in a significant regression (P=0.005) (Table 3 and Figure 4C).

Regression A is not statistically significant (P=0.163). There was no correlation between use and air concentration (Table 3, Figure 4A) and a township cap cannot be calculated.

Regression B also shows no statistical significance ($P=0.597$) (Table 3). Air concentrations detected in Ventura were significantly lower than other areas with comparable use regardless monitoring years (Figures 4A and 4B, additional analysis in Appendix 2). Ventura concentrations could be lower due to differences in application methods between the regions, weather, or other factors, such as geographic location facing the ocean and frequent strong wind blowing during application season. Minimal data size limits any firm conclusions for the low concentrations detected in Ventura.

Regression C is statistically significant ($P=0.005$) when excluding the data from Ventura County (Table 3, Figure 4C). The coefficient of determination (R^2) for this regression was the highest (0.61). It measured 61% of the air concentration (Y)'s variation that was explained by the monthly methyl bromide use (X).

Therefore the linear regression model C:

$$\text{Air concentration (ppb)} = 0.88 + 0.000024 * \text{township use (pounds/month)}$$

was used for further data analysis. Its residues were plotted in Figure 5 and coefficient estimates were listed in Table 4.

The Ryan-Joiner test, similar to Shapiro-Wilk test (Minitab 14), was performed to test normality of the regression residuals (Figure 6). The test failed to reject normality ($P>0.100$).

Homogeneity of variance was tested by the Levene test which looked for a linear relationship between the absolute values of the residual errors and the predictor variable (Willits, 2004). The tested linear relationship was not significant ($P=0.136$) as shown in Table 5. The test failed to establish heterodasticity.

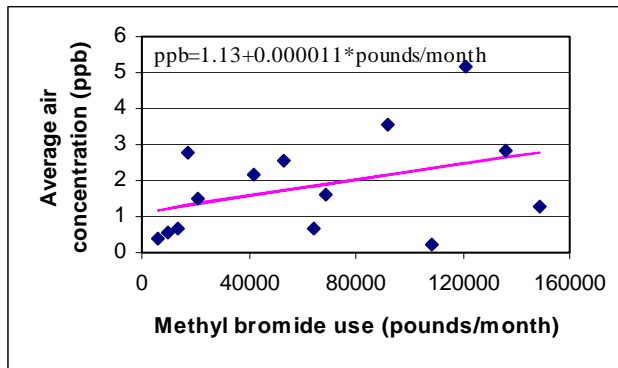
Due to the sample size, these tests are generally lacking in power to definitively reject normality or heterodasticity. At issue in the case of heterodasticity is whether the residuals show a fan-shaped pattern (residuals versus fitted values). If the large residual at $x=17396$ is removed, the plot suggests increasing variance with x (Figure 5). Thus while the technical requirements for regression are satisfied, the small sample size and a suggestive residual plot weaken confidence that the regression assumptions are satisfied.

Table 3. Regression Coefficient of determination (R^2) and mean square error (MSE)

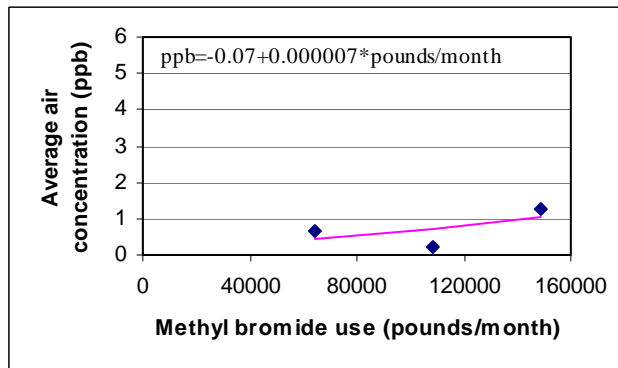
Data source	DF of residual error	R^2	MS		P
			regression	residual error	
A. all group data	12	0.16	4.027	1.826	0.163
B. only Ventura data	1	0.35	0.196	0.366	0.597
C. without Ventura data	9	0.61	12.46	0.897	0.005

Figure 4. Linear regression analyses using group data sets

A. all group data sets



B. only Ventura County data sets



C. excluding Ventura County data sets

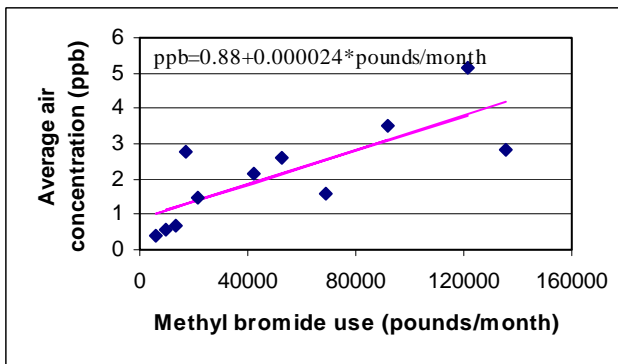


Table 4. Coefficient estimates for regression model C

	Coefficients	Standard Error	t Stat	P	Lower 95%	Upper 95%
Intercept	0.88	0.446	1.98	0.079	-0.13	1.89
Slope	0.000024	0.000006	3.73	0.005	0.000009	0.000039

Figure 5. Residual plots for the regression model C

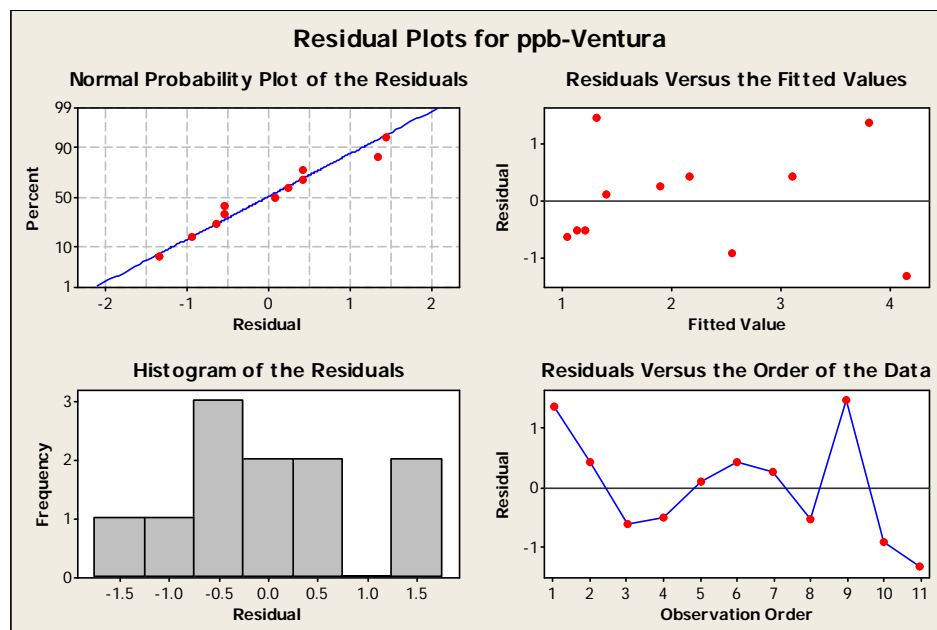


Figure 6. Ryan-Joiner Normality test for residues of regression model C

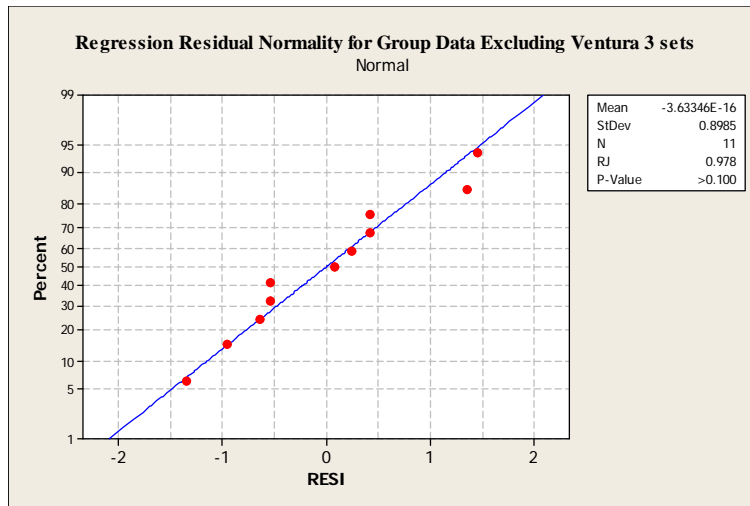


Table 5. ANOVA of the Levene test for linear regression residuals

Source	DF	SS	MS	F	P
Regression	1	0.521	0.521	2.6	0.136
Residual Error	9	1.748	0.194		
Total	10	2.269			

Methyl bromide monthly township use caps

The relationship between use and air concentrations for Ventura was different from the other areas (Figure 4). Air concentrations detected in Ventura (2002, 2005 and 2006) were significantly lower than other areas with comparable use levels. DPR staff are uncertain if the lower Ventura concentrations are due to differences in application methods between the regions or other factors. Ventura data were excluded from the regression used to determine township caps for several reasons. First, by excluding the Ventura data, the correlation between use and air concentration becomes statistically significant and a township cap can be calculated. Second, the data are more representative of the area affected by the township cap. Nine of the 11 data points represent monitoring in Monterey and Santa Cruz Counties. Most if not all townships affected by the cap are in Monterey and Santa Cruz Counties. Third, the slope of the regression line without Ventura data is steeper than the regression line including Ventura data. This results in more health-protective township caps for use levels greater than 60,000 pounds in a township. Fourth, there were several laboratory problems during the 2005 monitoring. Repeated instrument failures caused some samples to be held past the verified hold time, laboratory control samples were not included in the initial sets of samples because a gas standard was unavailable, and other

problems. Some of the 2005 data may have questionable validity. DPR staff have determined the recommended township cap from the regression that excludes Ventura data (model C). As specified in the risk management directive, the regulatory target level should be in the range of 2-9 ppb (Reardon 2009). Accordingly, the monthly township use caps can be derived based on regression model C (Table 6).

Table 6. Methyl bromide monthly township use caps

Ambient air concentration limits (ppb)	Township use derived by model (pounds/month)
2	46,667
3	88,333
4	130,000
5	171,667
6	213,333
7	255,000
8	296,667
9	338,333

Of the 202 California townships which had some methyl bromide use in 2008, DPR staff compiled the highest methyl bromide monthly use figures based on the 2008 Pesticide Use Report (Table 7). These are the townships that would most likely be affected by a cap on the township use. The township, county, reported use and corresponding regression model C concentration estimates are shown in Table 7.

Table 7. Highest monthly township methyl bromide use from 2008 Pesticide Use Report

Township	County	2008 Max Monthly Use (lbs)	Estimated Concentration (Model C - ppb)
M12S02E	MONTEREY/SANTA CRUZ	241,140	6.7
M14S02E	MONTEREY	196,620	5.6
M46N01W	SISKIYOU	153,084	4.6
M15S03E	MONTEREY	138,542	4.2
S10N34W	SANTA BARBARA	124,524	3.9
M12S01E	MONTEREY/SANTA CRUZ	79,650	2.8
S01N21W	VENTURA	79,080	2.8
S10N33W	SANTA BARBARA/SAN LUIS OBISPO	70,166	2.6
M14S03E	MONTEREY	67,083	2.5
S11N35W	SAN LUIS OBISPO	62,191	2.4
M47N01E	SISKIYOU	58,968	2.3
M11S02E	SANTA CRUZ	55,116	2.2
M14N03E	SUTTER	50,820	2.1
M20S17E	FRESNO	50,352	2.1

Conclusion

This data analysis found a linear regression model between monthly methyl bromide use and ambient air concentrations using aggregated data sets excluding data from Ventura County. The model weakly satisfied the regression assumptions of residual distribution normality and homogeneity of variance. The regression model was used to calculate monthly township caps based on various levels of concentration and to calculate monthly concentrations corresponding to 2008 monthly use in high methyl bromide use townships in California. The exclusion of Ventura data results in more health-protective township caps for use levels greater than 60,000 pounds in a township.

cc: Charles M. Andrews, DPR Associate Director
Marylou Verder-Carlos, DPR Assistant Director
John S. Sanders, Ph.D., DPR Environmental Program Manager
Susan Edmiston, DPR Branch Chief
Anna Fan-Cheuk, Office of Environmental Health Hazard Assessment Supervising Toxicologist

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Appendix 1
Statistical consultation summary from Willits (2009)

Here are the analyses I ran for you. I've attached four files. The first is a Word-formatted file containing the SAS output from those analyses. The others are probability plots for various forms of the residual errors from those analyses. The mixed model ANOVAs that I ran were primarily to see if the outliers in the analyses she had run were due to the fact that a few surprisingly large observations were also ones for which there was less replication (averaging over usage sites within a township) going on. These analyses included the township (called group) in the model and so if there was appreciable site to site variability within a township, it would "expect" an underreplicated observation to be more variable than one that was based on a greater number of replicate usage sites. In the first proc mixed, the estimated township to township variability was comparable in magnitude to the residual error, which means that those underreplicated townships would be seen as less extreme. In this analysis, there was a significant usage effect ($p = .0192$), which just means that there seems to be a non-zero relationship between usage and ambient concentrations.

The next part of the output (starting on SAS page 3 of the output -- look above the date to find these page numbers). Since there are two sources of randomness, there are several ways of calculating the residual errors, but all three of them are telling pretty much the same story. The outliers that came to Shifang's attention don't look that extreme in the residual plots. As such, it's hard to argue that they should be deleted because they're so extreme. (Mind you, I've said in the past that I don't like doing this, but even if I did, it would be hard to justify here.) At the same time, the residual distributions are significantly non-normal (all three have p-values in the range of .01 to .02) so the application of a normal tolerance interval to these data is hard to justify. There are nonparametric alternatives to parametric tolerance intervals, but they require a larger sample size than she has available.

The final part of the output (starting on SAS' page 12) contains the results of another mixed model analysis that looks at the question of whether the relationship between usage and ambient levels has changed since the latest set of regulations were implemented. The answer to this is "yes", as reflected by the significant interaction between the covariate (i.e., usage) and the ambient level ($p = .0336$).

If you have any questions about these results, please let me know.

-Neil Willits
Department of Statistics.

regression with random group effect

1
09:45

Thursday, December 10, 2009

The Mixed Procedure

Model Information

Data Set	WORK.SFAN
Dependent Variable	ppb
Covariance Structure	Variance Components
Estimation Method	REML
Residual Variance Method	Profile
Fixed Effects SE Method	Model-Based
Degrees of Freedom Method	Containment

Class Level Information

Class	Levels	Values
year	5	2000 2001 2002 2005 2006
group	9	1 2 3 4 5 6 7 8 9

Dimensions

Covariance Parameters	2
Columns in X	2
Columns in Z	9
Subjects	1
Max Obs Per Subject	34

Number of Observations

Number of Observations Read	34
Number of Observations Used	34
Number of Observations Not Used	0

Iteration History

Criterion	Iteration	Evaluations	-2 Res Log Like
	0	1	152.07188922
0.02861658	1	3	140.36597859

0.00287888	2	2	138.81819670
0.00047299	3	2	138.74834094
0.00000815	4	1	138.72794164
0.00000000	5	1	138.72761193

Convergence criteria met.
 regression with random group effect

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The Mixed Procedure

Covariance Parameter
 Estimates

Cov Parm	Estimate
group	1.2166
Residual	1.3064

Fit Statistics

-2 Res Log Likelihood	138.7
AIC (smaller is better)	142.7
AICC (smaller is better)	143.1
BIC (smaller is better)	143.1

Type 3 Tests of Fixed Effects

Effect	Num DF	Den DF	F Value	Pr > F
usage	1	24	6.30	0.0192
normality of residuals				

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Thursday, December 10, 2009

The UNIVARIATE Procedure
 Variable: Resid (Residual)

Moments

34	N	34	Sum Weights
0	Mean	0	Sum Observations
1.04454482	Std Deviation	1.02202975	Variance
2.83517555	Skewness	1.01723541	Kurtosis
34.469979	Uncorrected SS	34.469979	Corrected SS
0.17527666	Coeff Variation	.	Std Error Mean

Basic Statistical Measures

Location		Variability	
Mean	0.00000	Std Deviation	1.02203
Median	-0.06758	Variance	1.04454
Mode	.	Range	5.46902
		Interquartile Range	1.05097

Tests for Location: Mu0=0

Test	-Statistic-	-----p Value-----
Student's t	t 0	Pr > t 1.0000
Sign	M -3	Pr >= M 0.3915
Signed Rank	S -37.5	Pr >= S 0.5295

Tests for Normality

Test	--Statistic---	-----p Value-----
Shapiro-Wilk	W 0.919589	Pr < W 0.0157
Kolmogorov-Smirnov	D 0.154101	Pr > D 0.0398
Cramer-von Mises	W-Sq 0.169067	Pr > W-Sq 0.0132
Anderson-Darling	A-Sq 0.984658	Pr > A-Sq 0.0125

Quantiles (Definition 5)

Quantile	Estimate
100% Max	3.2194067
99%	3.2194067

95% 2.4258410
 90% 1.0809277
 75% Q3 0.3964934
 50% Median -0.0675824
 normality of residuals

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Thursday, December 10, 2009

The UNIVARIATE Procedure
 Variable: Resid (Residual)

Quantiles (Definition 5)

Quantile	Estimate
25% Q1	-0.6544774
10%	-0.9424248
5%	-1.4659792
1%	-2.2496092
0% Min	-2.2496092

Extreme Observations

-----Lowest-----		-----Highest-----	
Value	Obs	Value	Obs
-2.249609	14	0.71713	12
-1.465979	28	1.08093	15
-1.225760	21	1.48133	4
-0.942425	6	2.42584	10
-0.843297	25	3.21941	5

Stem Leaf	#	Boxplot
3 2	1	0
2		
2 4	1	0
1 5	1	
1 1	1	
0 677	3	
0 0112244	7	+-----+
-0 3332111110	10	*-----*
-0 9887776	7	+-----+
-1 2	1	
-1 5	1	
-2 2	1	0

-----+-----+-----+-----+

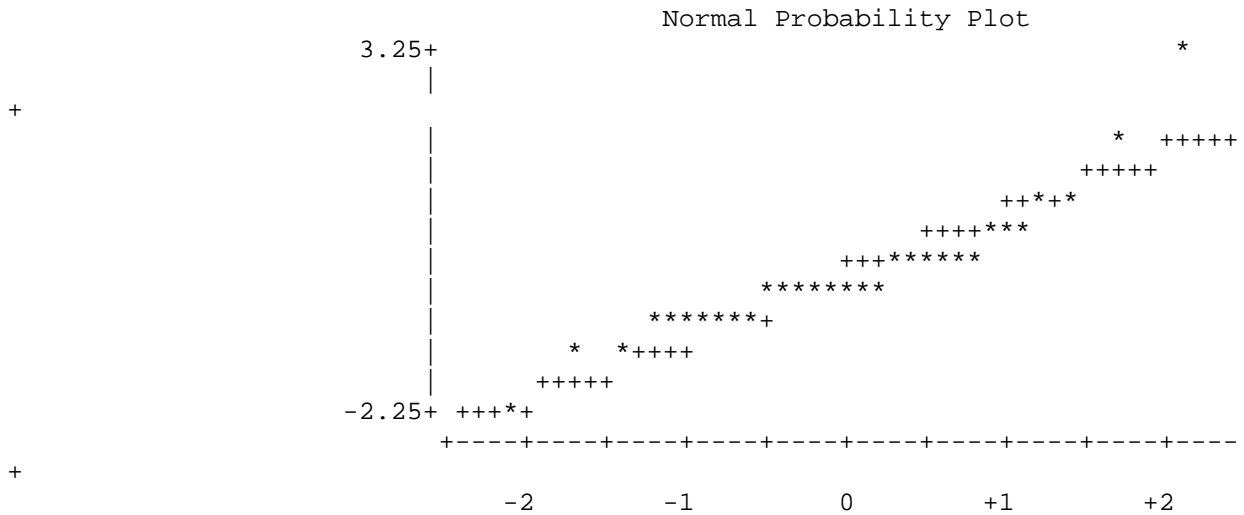
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normality of residuals

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The UNIVARIATE Procedure
Variable: Resid (Residual)



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normality of residuals

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The UNIVARIATE Procedure
Variable: StudentResid (Studentized Residual)

Moments

34	N	34	Sum Weights	
0.0713096	Mean	-0.0020973	Sum Observations	-
1.01459452	Std Deviation	1.00727083	Variance	
2.9236927	Skewness	1.05564029	Kurtosis	
33.4816191	Uncorrected SS	33.4817686	Corrected SS	

0.17274552 Coeff Variation -48026.071 Std Error Mean

Basic Statistical Measures

Location		Variability	
Mean	-0.00210	Std Deviation	1.00727
Median	-0.06235	Variance	1.01459
Mode	.	Range	5.40551
		Interquartile Range	0.99796

Tests for Location: Mu0=0

Test	-Statistic-	-----p Value-----	
Student's t	t -0.01214	Pr > t	0.9904
Sign	M -3	Pr >= M	0.3915
Signed Rank	S -39.5	Pr >= S	0.5076

Tests for Normality

Test	--Statistic--	-----p Value-----	
Shapiro-Wilk	W 0.919914	Pr < W	0.0161
Kolmogorov-Smirnov	D 0.150955	Pr > D	0.0476
Cramer-von Mises	W-Sq 0.164148	Pr > W-Sq	0.0157
Anderson-Darling	A-Sq 0.958119	Pr > A-Sq	0.0150

Quantiles (Definition 5)

Quantile	Estimate
100% Max	3.2237208
99%	3.2237208
95%	2.3287671
90%	1.0994548
75% Q3	0.3679625
50% Median	-0.0623469
normality of residuals	

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Thursday, December 10, 2009

The UNIVARIATE Procedure
 Variable: StudentResid (Studentized Residual)

Quantiles (Definition 5)

Quantile	Estimate
25% Q1	-0.6300004
10%	-0.9075477
5%	-1.4600988
1%	-2.1817892
0% Min	-2.1817892

Extreme Observations

-----Lowest-----		-----Highest-----	
Value	Obs	Value	Obs
-2.181789	14	0.70353	12
-1.460099	28	1.09945	15
-1.233748	21	1.45484	4
-0.907548	6	2.32877	10
-0.790778	25	3.22372	5

Stem Leaf	#	Boxplot
3 2	1	0
2		
2 3	1	0
1 5	1	
1 1	1	
0 677	3	
0 0112244	7	+-----+
-0 4332111110	10	*---*---
-0 9887766	7	+-----+
-1 2	1	
-1 5	1	
-2 2	1	0
-----+-----+-----+-----+		

normality of residuals

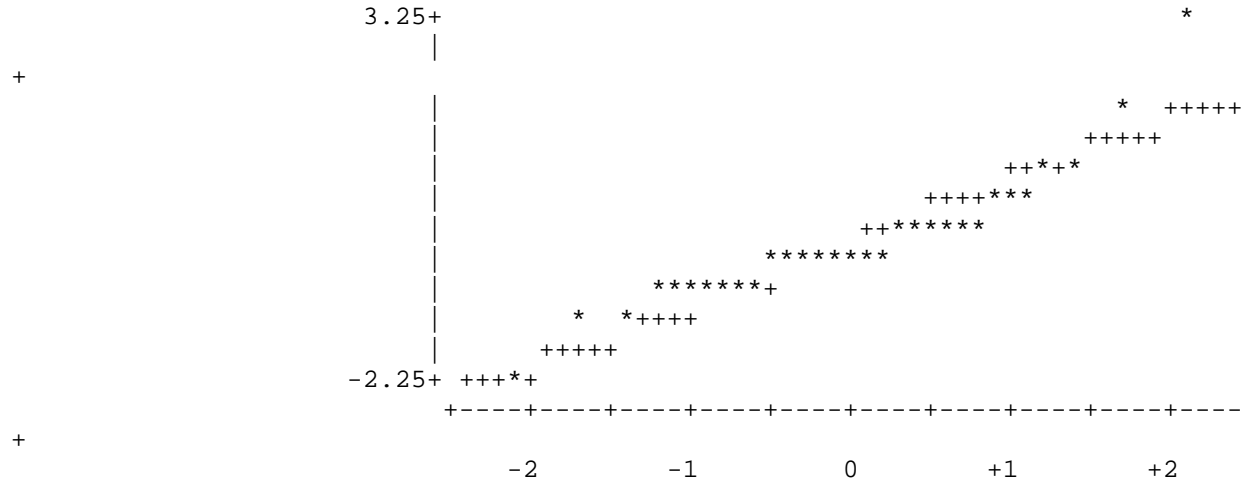
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The UNIVARIATE Procedure
 Variable: StudentResid (Studentized Residual)

Normal Probability Plot



normality of residuals

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Thursday, December 10, 2009

The UNIVARIATE Procedure
 Variable: PearsonResid (Pearson Residual)

Moments

	N	34	Sum Weights
Mean		0	Sum Observations
Std Deviation	0.8941795		Variance
Skewness	1.01723541		Kurtosis
Uncorrected SS	26.3853804		Corrected SS
Coeff Variation	.		Std Error Mean

Basic Statistical Measures

Location		Variability	
Mean	0.00000	Std Deviation	0.89418
Median	-0.05913	Variance	0.79956
Mode	.	Range	4.78487
		Interquartile Range	0.91950

Tests for Location: Mu0=0

Test	-Statistic-	-----p Value-----
Student's t	t 0	Pr > t 1.0000
Sign	M -3	Pr >= M 0.3915
Signed Rank	S -37.5	Pr >= S 0.5295

Tests for Normality

Test	--Statistic---	-----p Value-----
Shapiro-Wilk	W 0.919589	Pr < W 0.0157
Kolmogorov-Smirnov	D 0.154101	Pr > D 0.0398
Cramer-von Mises	W-Sq 0.169067	Pr > W-Sq 0.0132
Anderson-Darling	A-Sq 0.984658	Pr > A-Sq 0.0125

Quantiles (Definition 5)

Quantile	Estimate
100% Max	2.8166768
99%	2.8166768
95%	2.1223817
90%	0.9457096
75% Q3	0.3468942
50% Median	-0.0591282

normality of residuals

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Thursday, December 10, 2009

The UNIVARIATE Procedure
 Variable: PearsonResid (Pearson Residual)

Quantiles (Definition 5)

Quantile	Estimate
25% Q1	-0.5726059
10%	-0.8245327
5%	-1.2825934
1%	-1.9681956
0% Min	-1.9681956

Extreme Observations

-----Lowest-----		-----Highest-----	
Value	Obs	Value	Obs
-1.968196	14	0.627421	12
-1.282593	28	0.945710	15
-1.072425	21	1.296028	4
-0.824533	6	2.122382	10
-0.737805	25	2.816677	5

Stem Leaf	#	Boxplot
2 8	1	0
2 1	1	0
1		
1 3	1	
0 5669	4	
0 0111234	7	+---+---+
-0 3322111100	10	*-----*
-0 8776665	7	+---+---+
-1 31	2	
-1		
-2 0	1	0

-----+-----+-----+-----+

normality of residuals

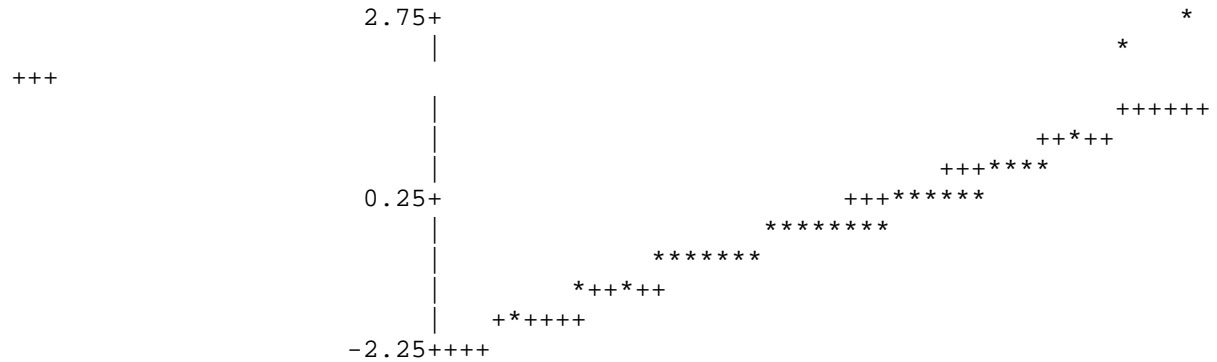
11

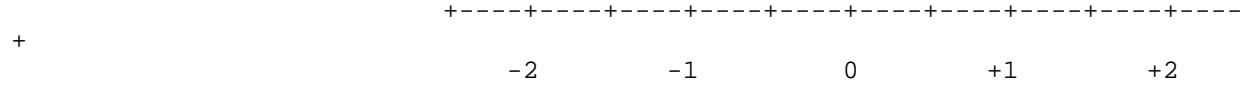
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The UNIVARIATE Procedure
 Variable: PearsonResid (Pearson Residual)

Normal Probability Plot





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ANCOVA for shifts in slope

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Thursday, December 10, 2009

The Mixed Procedure

Model Information

Data Set	WORK.RESIDUALS
Dependent Variable	ppb
Covariance Structure	Variance Components
Estimation Method	REML
Residual Variance Method	Profile
Fixed Effects SE Method	Model-Based
Degrees of Freedom Method	Containment

Class Level Information

Class	Levels	Values
year	5	2000 2001 2002 2005 2006
period	2	older recent
group	9	1 2 3 4 5 6 7 8 9

Dimensions

Covariance Parameters	2
Columns in X	6
Columns in Z	9
Subjects	1
Max Obs Per Subject	34

Number of Observations

Number of Observations Read	34
Number of Observations Used	34
Number of Observations Not Used	0

Iteration History

Criterion	Iteration	Evaluations	-2 Res Log Like
	0	1	160.17554669
0.00002232	1	3	151.23560173
0.00000003	2	1	151.23455688
0.00000000	3	1	151.23455535

Convergence criteria met.
 ANCOVA for shifts in slope

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Thursday, December 10, 2009

The Mixed Procedure

Covariance Parameter Estimates

Cov Parm	Estimate
group	1.2661
Residual	1.0885

Fit Statistics

-2 Res Log Likelihood	151.2
AIC (smaller is better)	155.2
AICC (smaller is better)	155.7
BIC (smaller is better)	155.6

Solution for Fixed Effects

Pr > t	Effect	period	Estimate	Standard Error	DF	t Value
0.6947	Intercept		0.5461	1.3348	7	0.41
0.8731	usage		-1.19E-6	7.375E-6	23	-0.16
0.9683	period	older	0.05892	1.4647	23	0.04

.	period	recent	0	.	.	.
0.0336	usage*period	older	0.000021	9.158E-6	23	2.26
.	usage*period	recent	0	.	.	.

Type 3 Tests of Fixed Effects

Effect	Num DF	Den DF	F Value	Pr > F
usage	1	23	4.00	0.0575
period	1	23	0.00	0.9683
usage*period	1	23	5.11	0.0336

Appendix 2. Comparison of data from Ventura to data from other counties

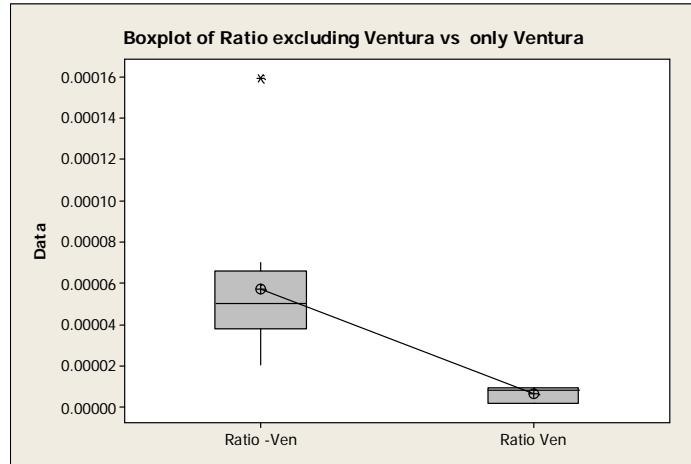
Data from Ventura County were different from other counties regardless monitoring year. The three group data sets, one was in 2002 monitored by AMBI and the other two were in 2005 and 2006 monitored by ARB, generally showed high use but low air concentrations.

To statistically compare data from Ventura to those from other counties, the ratio of the air concentration to the normalized monthly township use was calculated for each group (Table 1). Student t-test was performed to compare the mean of the ratios for data excluding Ventura County vs. those of only Ventura County. The results showed significant difference at P=0.04 (Table 1). Figure 1 is a boxplot of this comparison.

Table 1. t-test for the ratio of air concentration/monthly township use using group data sets

	Excluding Ventura		Only Ventura
	County	Ratio	Ratio
Ratio of concentration/use (ppb/pounds per month)	Santa Cruz	4.26E-05	8.68E-06
	Santa Cruz	3.83E-05	2.21E-06
	Monterey	4.89E-05	9.94E-06
	Monterey	5.11E-05	
	Monterey	6.61E-05	
	Monterey	5.03E-05	
	Monterey	5.87E-05	
	Kern	6.99E-05	
	Kern	1.59E-04	
	Monterey	2.32E-05	
	Santa Cruz	2.08E-05	
N	11		3
Mean	5.72E-05		6.95E-06
Standard deviation	3.72E-05		4.15E-06
SE mean	1.10E-05		2.40E-06
DF		12	
T-value		2.27	
P-value		0.04	

Figure 1. Boxplot comparison of excluding Ventura versus Ventura only



Further analysis of One Way ANOVA compared the ratio of the grouped air concentration to the grouped normalized monthly township use by counties and years (Figure 2), respectively (The data were unbalanced, so not suitable for Two Way ANOVA). The results showed a significant difference between counties ($P=0.003$), and the result of residual normality test was $P=0.07$. However, there was no significant difference among years ($P=0.143$), and the residues was not normally distributed ($p<0.01$). Figure 2 showed that the ratio of the air concentration to the normalized monthly township use for Ventura County was much lower than all other counties.

Figure 2. One Way ANOVA for the ratio of the air concentration to the normalized monthly township use:

